Separating sparse signals from correlated noise in binary classification

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(e.g. a disease)

Statistical genetics: predict the effect of genes on observable outcomes



genotypes correlate with geographic location





(e.g. a disease)



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Statistical genetics: predict the effect of genes on observable outcomes



genotype

genotypes correlate with geographic location





Problem:

indirect associations "population structure" **Goal:** separate these effects

cultural factors climate factors economic conditions political factors

causal association

 $\{-1,+1\}$

binary phenotype (e.g. a disease)



Novembre et al, Nature 2008









Confounding by population structure



- Population structure: common cause of genes and phenotypes
- Goal: estimate causal effect
- Problem: generative mechanism is unknown



Linear Mixed Models for Regression

- Linear mixed model (LMM): widely appreciated in genetics
- Linear regression + correlated noise.

 $X \in \mathbb{R}^{d imes n}: ext{genotypes} \ y \in \mathbb{R}^n: ext{phenotypes} \ w \in \mathbb{R}^d: ext{weight vector}$

$$y_i = X_i^ op w + \epsilon_i ~~ \epsilon \sim \mathcal{N}(0,\Sigma)$$



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 ,

Multivariate noise:

- Allows to express similarities between samples
- Typical choice in genetics: $\Sigma = \lambda_0 \mathbf{I} + \lambda_1 X^{ op} X$



Equivalent model formulation with a linear kernel:

$$y = X^ op w + X^ op w' + \epsilon, \quad w' \sim \mathcal{N}(0,\lambda_1\mathbf{I}), \quad \epsilon \sim \mathcal{N}(0,\lambda_0\mathbf{I})$$

We see that w and w' play similar roles, but:

- w is assumed to be a large, sparse vector (causal)
- w' is dense and unobserved (confounder)



A solution can be obtained by transforming X and y:

$$egin{aligned} \Sigma &= BB^ op, \quad y' = B^{-1}y, \quad X'^ op = B^{-1}X^ op \ y' &= X'^ op w + \epsilon, \quad \epsilon \sim \mathcal{N}(0,I) \end{aligned}$$

- This results in a standard linear regression problem
- O(n^3) scaling if done naively
- State of the art for many applications in biology



Goal: generalize the LMM paradigm to classification **Idea:** probit regression model with correlated noise:

$$y_i = ext{sign}(X_i^ op w + \epsilon_i), \quad \epsilon \sim \mathcal{N}(0, \Sigma)$$

Challenge: exact inference becomes intractable due to the nonlinearity Solution: approximate inference (this talk)



For simplicity, assume $\forall_i: y_i = 1$ Likelihood that all examples are correctly classified:

$$\mathbb{P}ig(orall_i: y_i = ext{sign}(X_i^ op w + \epsilon_i)ig) = \int_{\mathbb{R}^n_+} \mathcal{N}(\epsilon; X^ op w, \Sigma) d^n \epsilon$$





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Objective function: negative log likelihood + regularizer. Define $\mu(w) = X^{ op} w$.

$$\mathcal{L}(w) = -\log \int_{\mathbb{R}^{n}_{+}} \mathcal{N}(\epsilon; \mu(w), \Sigma) d^{n}\epsilon + \underbrace{\lambda_{0} ||w||_{1}^{1}}_{=:\mathcal{L}^{\mathrm{reg}}(w)}$$
$$\underbrace{-:\mathcal{L}^{\mathrm{reg}}(w)}_{=:\mathcal{L}^{\mathrm{loss}}(w)}$$

Lasso regularizer: Favors sparsity



Minimizing the loss function



Minimizing the objective function leads to two computational problems:

- (i) intractable high-dimensional integral
- (ii) the l_1 -norm regularizer is not everywhere differentiable



Minimizing the loss function



Minimizing the objective function leads to two computational problems:

(i) intractable high-dimensional integral Solution: Expectation Propagation (EP)

(ii) the l_1 -norm regularizer is not everywhere differentiable



Minimizing the loss function



Minimizing the objective function leads to two computational problems:

(i) intractable high-dimensional integral Solution: Expectation Propagation (EP)

(ii) the l₁-norm regularizer is not everywhere differentiable Solution: Alternating Direction Method of Multipliers (ADMM)





$$\mathcal{L}(w) = \underbrace{-\log \int_{\mathbb{R}^{n}_{+}} \mathcal{N}(\epsilon; \mu(w), \Sigma) d^{n} \epsilon}_{=:\mathcal{L}^{\mathrm{loss}}(w)} + \underbrace{\lambda_{0} ||w||_{1}^{1}}_{=:\mathcal{L}^{\mathrm{reg}}(w)}$$



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$$p(\epsilon|\mu, \Sigma) = \frac{\mathbb{1}[\epsilon \in \mathbb{R}^n_+] \mathcal{N}(\epsilon; \mu, \Sigma)}{\int_{\mathbb{R}^n_+} \mathcal{N}(\epsilon; \mu, \Sigma) d^n \epsilon}$$





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$$\mu_p(w) = \mathbb{E}_{p(\epsilon|\mu(w),\Sigma)} [\epsilon],$$

$$\Sigma_p(w) = \mathbb{E}_{p(\epsilon|\mu(w),\Sigma)} [(\epsilon - \mu_p(w))(\epsilon - \mu_p(w))^\top]$$

$$\Delta \mu = \mu_p - \mu$$

$$\nabla_{w} \mathcal{L}^{\text{loss}}(w) = \Delta \mu \Sigma^{-1} X^{\top},$$

$$H^{\text{loss}}(w) = -X [\Sigma^{-1} (\Sigma_{p} - \Delta \mu \Delta \mu^{\top}) \Sigma^{-1} - \Sigma^{-1}] X^{\top}$$



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$$\mu_{p}(w) = \mathbb{E}_{p(\epsilon|\mu(w), \Sigma)} [\epsilon],$$

$$\sum_{p(w)} = \mathbb{E}_{p(\epsilon|\mu(w), \Sigma)} [\epsilon],$$

$$\Delta \mu = \mu_{p} - \mu$$

$$\sum_{p(w)} \int_{\mathbb{R}^{n}_{+}} \mathcal{N}(e; \mu, \Sigma) d^{n} \epsilon$$

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$$\sum_{p(w)} \int_{\mathbb{R}^{n}_{+}} \mathcal{N}(e; \mu, \Sigma) [\epsilon] + \frac{1}{(\epsilon - \mu_{p}(w))(\epsilon - \mu_{p}(w))^{\top}} = \frac{1}{(\epsilon - \mu_{p}(w))(\epsilon - \mu_{p}(w))^{\top}}$$

$$\sum_{p(w)} \int_{\mathbb{R}^{n}_{+}} \mathcal{N}(e; \mu, \Sigma) [\epsilon] + \frac{1}{(\epsilon - \mu_{p}(w))(\epsilon - \mu_{p}(w))^{\top}} = \frac{1}{(\epsilon - \mu_{p}(w))(\epsilon - \mu_{p}(w))^{\top}} = \frac{1}{(\epsilon - \mu_{p}(w))(\epsilon - \mu_{p}(w))^{\top}}$$

$$V_w \mathcal{L}^{-}(w) = \Delta \mu \Delta^{-} X^{-},$$

$$H^{\text{loss}}(w) = -X [\Sigma^{-1} (\Sigma_p - \Delta \mu \Delta \mu^{\top}) \Sigma^{-1} - \Sigma^{-1}] X^{\top}$$



We use **Expectation Propagation** to approximate $p(\epsilon|\mu, \Sigma)$ by a Gaussian $q(\epsilon; \mu_q, \Sigma_q) = \mathcal{N}(\epsilon; \mu_q, \Sigma_q)$ Then: $\mu_p \approx \mu_q \quad \Sigma_p \approx \Sigma_q$



- The correlated probit model can be seen as a generalization of various other models.
- We compared the performance gain over these methods in our experiments.

 $\begin{array}{ll} \text{Uncorrelated Probit:} & y = \operatorname{sign}(X^\top w + \epsilon), \quad \epsilon \sim \mathcal{N}(0, \mathrm{I}), \quad w \sim \operatorname{Laplace}(.\,;\lambda_0) \\ \\ \text{GP classification:} & y = \operatorname{sign}(f), \quad f \sim \mathcal{N}(0, \Sigma(X)) \\ \\ \text{LMM Lasso:} & y = X^\top w + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \Sigma), \quad w \sim \operatorname{Laplace}(.\,;\lambda_0) \end{array}$



Experiments (Simulated Data)

- Generate artificial data from the model
- Varied the amount of non-zero weights w_i
- Compute accuracies for different levels of sparsity





Experiments (TBC)

• Predict Tuberculosis based on gene expression levels.

Tuberculosis data set: Berry et. al., Nature 466, 2010.

• Confounding by populations structure.





Experiments (TBC)

- Tuberculosis data set
- Correlate w with largest eigenvalue of Σ

Tuberculosis data set: Berry et. al., Nature 466, 2010.





- Algorithm for sparse feature selection in binary classification, where the data are confounded
- Signals found by our model are less correlated with the confounders
- Improved prediction performances
- Future: employ scalable MCMC to sample from the posterior
- Data subsampling is possible
- problem: high-dimensional feature space dimensionality d.



Thank you.









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